



Moments of Markov-switching models

Anna Petričková
ISCAMI 2012
Malenovice
12.5.2012



Introduction

- Overview to MSW models
- Moments of MSW models
 - State-independent MSW model
 - State-dependent MSW model



Overview to MSW models

- If we come out from AR(p), then
$$X_t = \phi_{0,q_t} + \phi_{1,q_t} X_{t-1} + \dots + \phi_{p,q_t} X_{t-p} + \varepsilon_t \quad \text{for } q_t = 1, 2$$
where $\{\varepsilon_t\}$ - process i.i.d. $\approx N(0, 1)$.
- special case of discrete first order Markov process (important is only the actual and the previous state)
=> regime q_t depends only on the regime q_{t-1}
$$P(q_t = S_j \mid q_{t-1} = S_i) = p_{ij}$$
- state transition probabilities from one state to another:
$$\sum_{j=1}^N p_{ij} = 1 \quad p_{ij} \geq 0$$



Overview to MSW models

- **B** – the ($N \times N$) matrix of transition probabilities for the 'time-reversed' Markov chain that moves back in time

$$P(q_t = S_j \mid q_{t+1} = S_i) = b_{ij} \quad 0 \leq b_{ij} \leq 1 \quad \sum_{j=1}^N b_{ij} = 1$$

Since

$$\begin{aligned} P(q_t = S_j \cap q_{i+1} = S_i) &= P(q_{i+1} = S_i \mid q_t = S_j) P(q_t = S_j) = \\ &= P(q_t = S_j \mid q_{i+1} = S_i) P(q_{i+1} = S_i) \end{aligned}$$

the 'backward' transition probability matrix **B** is related to the 'forward' transition probabilities as follows:

$$b_{ij} = p_{ji} \left(\frac{\pi_j}{\pi_i} \right)$$



First order MSW model

a) Simple autoregressive MSW model

$$X_t = \phi_{q_t} + \phi_1 X_{t-1} + \varepsilon_t$$

b) State-dependent autoregressive MSW

$$X_t = \phi_{q_t} + \phi_{q_t,1} X_{t-1} + \varepsilon_t$$



First order MSW model

Let's have simple autoregressive MSW model

$$X_t = \phi_{q_t} + \phi_1 X_{t-1} + \varepsilon_t$$

then for state-dependent mean value holds

$$\mu_{q_t} = \phi_{q_t} + \phi_1 \mu_{q_{t-1}}$$

After substituting into the basic model we get

$$X_t = \mu_{q_t} + \phi_1 (X_{t-1} - \mu_{q_{t-1}}) + \varepsilon_t \quad (1)$$

and after substituting backwards we obtain

$$X_t - \mu_{q_t} = \sum_{i=0}^{\infty} \phi_1^i \varepsilon_{t-i}$$

Derivation of the second central moment

$$\begin{aligned} E[(X_t - \mu)^2] &= E[(\mu_{q_t} - \mu)^2 + \phi_1^2 (X_{t-1} - \mu_{q_{t-1}})^2 + \varepsilon_t^2] + \\ &\quad + 2\phi_1 \text{Cov}(\mu_{q_t} - \mu, X_{t-1} - \mu_{q_{t-1}}) + 2\text{Cov}(\mu_{q_t} - \mu, \varepsilon_t) + \\ &\quad + 2\phi_1 \text{Cov}(X_{t-1} - \mu_{q_{t-1}}, \varepsilon_t) \end{aligned}$$

- From the assumption $\{\varepsilon_t\}$ - process i.i.d. $\approx N(0, 1)$ and from independence between ε_{t-1} and q_t for $i = 0, \pm 1, \pm 2, \dots \Rightarrow$ all covariances = 0 \Rightarrow

$$\begin{aligned} E[(X_t - \mu)^2] &= E[E[(X_t - \mu)^2] | q_t] = \\ &= E[E[(\mu_{q_t} - \mu)^2 + \phi_1^2 (X_{t-1} - \mu_{q_{t-1}})^2 + \varepsilon_t^2] | q_t] = \end{aligned}$$

Derivation of the second central moment

$$= \pi' \left(\begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \otimes \begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \right) + \pi' \phi_1^2 \left(\sum_{i=0}^{\infty} \phi_1^{2i} \mathbf{P}^i \begin{pmatrix} \sigma_1^2 \\ \sigma_2^2 \\ \sigma_3^2 \end{pmatrix} \right) + \pi' \begin{pmatrix} \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \end{pmatrix}$$

$$= \pi' \left(\begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \otimes \begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \right) + \phi_1^2 \left(\sum_{i=0}^{\infty} \phi_1^{2i} \pi' \mathbf{P}^i \begin{pmatrix} \sigma_1^2 \\ \sigma_2^2 \\ \sigma_3^2 \end{pmatrix} \right) + \pi' \begin{pmatrix} \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \end{pmatrix} =$$

$$= \pi' \left(\begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \otimes \begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \right) + \phi_1^2 \left(\sum_{i=0}^{\infty} \phi_1^{2i} \pi' \begin{pmatrix} \sigma_1^2 \\ \sigma_2^2 \\ \sigma_3^2 \end{pmatrix} \right) + \pi' \begin{pmatrix} \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \end{pmatrix} =$$

Derivation of the second central moment

$$\begin{aligned}
 &= \pi' \left[\begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \otimes \begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} + \phi_1^2 (1 - \phi_1^2)^{-1} \begin{pmatrix} \sigma_1^2 \\ \sigma_2^2 \\ \sigma_3^2 \end{pmatrix} + \begin{pmatrix} \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \end{pmatrix} \right] = \\
 &= \pi' \left[\begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \otimes \begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} + \begin{pmatrix} \frac{\phi_1^2 \sigma_1^2}{(1 - \phi_1^2)} \\ \frac{\phi_1^2 \sigma_2^2}{(1 - \phi_1^2)} \\ \frac{\phi_1^2 \sigma_3^2}{(1 - \phi_1^2)} \end{pmatrix} + \begin{pmatrix} \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \end{pmatrix} \right]
 \end{aligned}$$

- In the proof we use that $\sum_{i=0}^{\infty} \phi^{2i} = \frac{1}{1 - \phi^2}$ and π' is vector of steady-state probabilities, where $\pi' \mathbf{P} = \pi'$.

The simple state-independent first-order autoregressive MSW model

2. central moment

$$E[(y_t - \mu)^2] = \pi' \left[\begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \otimes \begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} + \begin{pmatrix} \frac{\phi_1^2 \sigma_1^2}{(1 - \phi_1^2)} \\ \frac{\phi_1^2 \sigma_2^2}{(1 - \phi_1^2)} \\ \frac{\phi_1^2 \sigma_3^2}{(1 - \phi_1^2)} \end{pmatrix} + \begin{pmatrix} \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \end{pmatrix} \right]$$



The second central moment of MSW model for 3 states (case **b**)

From

$$X_t = \mu_{q_t} + \phi_{1,q_{t-1}}(X_{t-1} - \mu_{q_{t-1}}) + \varepsilon_t$$

where $\phi_{1,q_{t-1}}$ is first-order autoregressive coefficient for state q_{t-1} . After substituting backwards we get

$$X_t - \mu_{q_t} = \sum_{l=1}^{\infty} \left(\prod_{j=1}^l \phi_{1,q_{t-j}} \right) \varepsilon_{t-j} + \varepsilon_t$$

Derivation of the second central moment

$$\begin{aligned}
 E[(X_t - \mu)^2] = & E[(\mu_{q_t} - \mu)^2 + \phi_{1,q_{t-1}}^2 (X_{t-1} - \mu_{q_{t-1}})^2 + \varepsilon_t^2] + \\
 & + 2\text{Cov}(\mu_{q_t} - \mu, \varepsilon_t) + 2\text{Cov}(\mu_{q_t} - \mu, \phi_{1,q_{t-1}} (X_{t-1} - \mu_{q_{t-1}})) + \\
 & + 2\text{Cov}(\phi_{1,q_{t-1}} (X_{t-1} - \mu_{q_{t-1}}), \varepsilon_t)
 \end{aligned}$$

From the assumption $\{\varepsilon_t\}$ - process i.i.d. $\approx N(0, 1)$ and from independence between ε_{t-1} and q_t for $i = 0, \pm 1, \pm 2, \dots \Rightarrow$ all covariances = 0 \Rightarrow

$$E[(X_t - \mu)^2] = E[(\mu_{q_t} - \mu)^2 + \phi_{1,q_{t-1}}^2 (X_{t-1} - \mu_{q_{t-1}})^2 + \varepsilon_t^2]$$

$$E[(\mathbf{X}_t - \boldsymbol{\mu}_{\mathbf{q}_t})^2 | \mathbf{q}_t] = E[(\phi_{1,q_{t-1}}^2 (\mathbf{X}_{t-1} - \boldsymbol{\mu}_{\mathbf{q}_{t-1}})^2 + 2\phi_{1,q_{t-1}} (\mathbf{X}_{t-1} - \boldsymbol{\mu}_{\mathbf{q}_{t-1}}) \varepsilon_t + \varepsilon_t^2) | \mathbf{q}_t]$$

Derivation of the second central moment

$$E\left[\left(X_t - \mu_{q_t}\right)^2 | q_t = i\right] = \sum_{j=1}^3 E\left[\phi_{1,q_{t-1}}^2 \left(X_{t-1} - \mu_{q_{t-1}}\right)^2 | q_{t-1} = j \cap q_t = i\right] \cdot P(q_{t-1} = j | q_t = i) + \sigma_{\varepsilon}^2 =$$

$$= \sum_{j=1}^3 E\left[\phi_{1,q_{t-1}}^2 \left(X_{t-1} - \mu_{q_{t-1}}\right)^2 | q_{t-1} = j\right] \cdot b_{ij} + \sigma_{\varepsilon}^2$$

Label $\sigma_{\varepsilon}^2 = \begin{pmatrix} \sigma_{\varepsilon}^2 \\ \sigma_{\varepsilon}^2 \\ \sigma_{\varepsilon}^2 \end{pmatrix}$, Φ is 3 x 3 diagonal matrix of autoregressive coefficients $(\phi)_i, i=1,2,3$

Thus

$$E\left[\left(X_t - \mu_{q_t}\right)^2 | \mathbf{q}_t\right] = \mathbf{B} \cdot \Phi^2 E\left[\left(X_{t-1} - \mu_{q_{t-1}}\right)^2 | \mathbf{q}_{t-1}\right] + \sigma_{\varepsilon}^2$$

Derivation of the second central moment

We assume stationary process $\{X_t\} \Rightarrow$

$$E \begin{bmatrix} X_t | q_t = 1 \\ X_t | q_t = 2 \\ X_t | q_t = 3 \end{bmatrix} = E \begin{bmatrix} X_{t-1} | q_{t-1} = 1 \\ X_{t-1} | q_{t-1} = 2 \\ X_{t-1} | q_{t-1} = 3 \end{bmatrix} = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \end{pmatrix}$$

$$E \left[(X_{t-1} - \mu_{q_{t-1}})^2 | q_{t-1} \right] = (\mathbf{I}_3 - \mathbf{B} \cdot \Phi^2)^{-1} \sigma_\varepsilon^2$$

and thus

$$\begin{aligned} E \left[(X_t - \mu)^2 \right] &= E \left[E \left[(X_t - \mu)^2 \right] | q_t \right] = \\ &= E \left[E \left[\left(\mu_{q_t} - \mu \right)^2 + \phi_{1,q_{t-1}}^2 (X_{t-1} - \mu_{q_{t-1}})^2 + \varepsilon_t^2 \right] | q_t \right] = \\ &= \pi' \begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} \otimes \begin{pmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{pmatrix} + \pi' \Phi^2 (\mathbf{I}_3 - \mathbf{B} \Phi^2)^{-1} \begin{pmatrix} \sigma_1^2 \\ \sigma_2^2 \\ \sigma_3^2 \end{pmatrix} + \pi' \begin{pmatrix} \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \end{pmatrix} \end{aligned}$$

State-dependent first-order autoregressive MSW model

2. central moment

$$E[(X_t - \mu)^2] = \boldsymbol{\pi}' \begin{bmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{bmatrix} \otimes \begin{bmatrix} \mu_1 - \mu \\ \mu_2 - \mu \\ \mu_3 - \mu \end{bmatrix} + \Phi^2 (\mathbf{I}_3 - \mathbf{B}\Phi^2)^{-1} \begin{bmatrix} \sigma_1^2 \\ \sigma_2^2 \\ \sigma_3^2 \end{bmatrix} + \begin{bmatrix} \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \\ \sigma_\varepsilon^2 \end{bmatrix}$$



Thank you for your attention